# Post-Construction Monitoring Study for the Headwaters Wind Farm Randolph County, Indiana

Year 4 Final Report April 1 – July 31, 2022



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# EXECUTIVE SUMMARY

Headwaters Wind Farm, LLC (Headwaters) is operating the Headwaters Wind Farm (Project) in Randolph County, Indiana. This report details the post-construction monitoring studies conducted in 2022, consistent with the Project's Habitat Conservation Plan (HCP) and Incidental Take Permit (ITP; TE85617C-0) for Indiana and northern long-eared bats (Covered Species). The Project obtained the ITP on June 4, 2019, and has since completed one fall season (2019) and two full-seasons (April 1 – October 15, 2020 and 2021) studies of monitoring. Monitoring was conducted in spring and summer of 2022 to complete the initial three full years of compliance monitoring. In 2022, turbines were operated with increased cut-in speeds during spring migration; turbines with designated summer risk per the Project's HCP were also operated with increased cut-in speeds during the summer.

Post-construction monitoring was completed in accordance with the study plan, which was approved by US Fish and Wildlife Service on February 23, 2022. The study plan was designed to achieve a 25% probability of detecting a single bat carcass (*g* of 0.25) for the 100 wind turbines at the Project. The overall goal of this post-construction fatality monitoring study was to generate reliable fatality estimates for the Covered Species and to evaluate compliance with the incidental take authorization granted under the Project's ITP. More specifically, the objectives of this study were to estimate take for the Covered Species using the Evidence of Absence (EoA) framework as outlined in the HCP and to determine if adaptive management was necessary to maintain compliance with the Project's ITP.

Standardized carcass searches for bat carcasses were completed at two plot types: full plots and road and pad plots, and were conducted by two types of searchers: technician and dog-handler team (consisting of one dog trained to detect carcasses and one handler). The frequency of searches varied across seasons, with more searches occurring when take of Covered Species was considered more likely to occur. Searcher efficiency and carcass persistence trials were also conducted during each season to correct for detection and scavenger bias.

No Covered Species were found at the Project. Sixty bats were found during the study. The most commonly found bat species were eastern red bat (46.7%), silver-haired bat (31.7%), hoary bat (11.7%), and big brown bat (10.0%). Species composition recorded at the Project was similar to previous studies at the Project and other wind facilities in Indiana. Fifteen bird carcasses were recorded; no federally or state-listed birds were found.

The *g* was 0.22 (90% confidence interval: 0.19-0.25). Based on data collected to date (2019, 2020, 2021 and 2022 studies), the EoA model estimated the mean annual fatality rates were 9.61 Indiana bats and 0.64 northern long-eared bats with an overall *g* of 0.26 (90% confidence interval: 0.24-0.28) throughout the initial three year monitoring period. The probability that the annual take rate exceeded the expected annual take rate was 0.46 for Indiana bat and 0.05 for northern long-eared bat. The cumulative take estimates through 2022 were 27 Indiana bat fatalities and zero northern long-eared bat fatalities. The estimated levels of Indiana bat and northern long-eared

bat take were below levels authorized within the ITP. No adaptive management actions are necessary at this time.

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#### **REPORT REFERENCE**

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# INTRODUCTION

Headwaters Wind Farm, LLC (Headwaters), a subsidiary of EDP Renewables North America, LLC (EDPR), is operating the Headwaters Wind Farm (Project) in Randolph County, Indiana. EDPR obtained an Incidental Take Permit (ITP; TE85617C-0, dated June 4, 2019) for the federally listed endangered Indiana bat (*Myotis sodalis*) and endangered<sup>1</sup> northern long-eared bat (*M. septentrionalis*; hereafter Covered Species) from the US Fish and Wildlife Service (USFWS). The Project had completed one fall-only season (2019) and two full-seasons (April 1 – October 15, 2020 and 2021) of monitoring prior. Monitoring was conducted in spring and summer of 2022 to complete three full years of initial compliance monitoring. This report presents the results of the fourth consecutive survey period of compliance monitoring conducted under the ITP from April 1, 2022 to July 31, 2022. The objectives of this study were to estimate take of the Covered Species using the Evidence of Absence (EoA) framework as outlined in the Habitat Conservation Plan (HCP; Headwaters 2019) and determine if adaptive management is necessary to maintain compliance with the Project's ITP.

# STUDY AREA

The primary land cover type within 100 meters (m; 328 feet [ft]) of the turbines (i.e., within the Permit Area) is cultivated crops, which covers 87.7% of the Permit Area (National Land Cover Database {NLCD] 2019). The next most common land cover is developed area (e.g., farmsteads) that composes approximately 5.5% of the site, followed by patches of deciduous forest composing 5.1% of the site and hay/pasture composing 1.1% (NLCD 2019). All other land cover types collectively make up less than one percent of the total land cover (Table 1, Figure 1; NLCD 2019).

The Project became fully operational in 2014, and consists of 100 2.0-megawatt (MW) Vestas V110 wind turbines that have a 95-meter (m; 311-ft) hub height and a 55-m (180-ft) blade length. All turbines are within the migratory range of the Covered Species, and EDPR adjusted turbine operations during the spring migration period to minimize impacts to the Covered Species. Additionally, EDPR adjusted turbine operations for 11 designated summer risk turbines during summer (Table 2).

<sup>&</sup>lt;sup>1</sup> The northern long-eared bat was listed as threatened when the ITP was received. Its status will change to endangered as of January 30, 2023.

Habitat	Acres	% Composition
Cultivated Crops	25,668	87.7
Developed <sup>1</sup>	1,602	5.5
Deciduous Forest	1,488	5.1
Hay/Pasture	320	1.1
Woody Wetlands	113	0.4
Herbaceous	55	0.2
Open Water	10	< 0.1
Emergent Herbaceous Wetlands	6	< 0.1
Mixed Forest	3	< 0.1
Evergreen Forest	2	< 0.1
Barren Land	1	< 0.1
Total <sup>2</sup>	29,269	100

 Table 1.
 Land cover types and percent composition at the Headwaters Wind Farm Permit

 Area, in Randolph County, Indiana.

<sup>1</sup> Developed areas include high-, medium-, and low-intensity developed areas, as well as developed open space.

<sup>2</sup> Sums may not equal total values shown due to rounding.

Data from National Land Cover Database. 2019

#### Table 2. Seasonal turbine operational regime at the Headwaters Wind Farm, Randolph County, Indiana, from April 1 – July 31, 2022.

				Feathering	
Season	Turbines	Time of Day	Cut-In Speed	Below Cut-In <sup>1</sup> ?	Temperature Threshold <sup>2</sup>
Spring (April 1 – May 15)	100 (All)	0.5 hour before sunset to 0.5 hour after sunrise	3.5 m/s	Yes	10 °C
Summer	89 turbines without summer risk	0.5 hour before sunset to 0.5 hour after sunrise	3.0 m/s	Yes	No
(May 16 – July 31)	11 turbines with summer risk	0.5 hour before sunset to 0.5 hour after sunrise	5.0 m/s	Yes	No

<sup>1</sup> Feathering means that turbine blades will be pitched into the wind such that they spin at less than one rotation per minute.

<sup>2</sup> Turbines will be feathered below cut-in when temperatures are above the threshold. In practice, the Project feathered on all nights regardless of temperature.

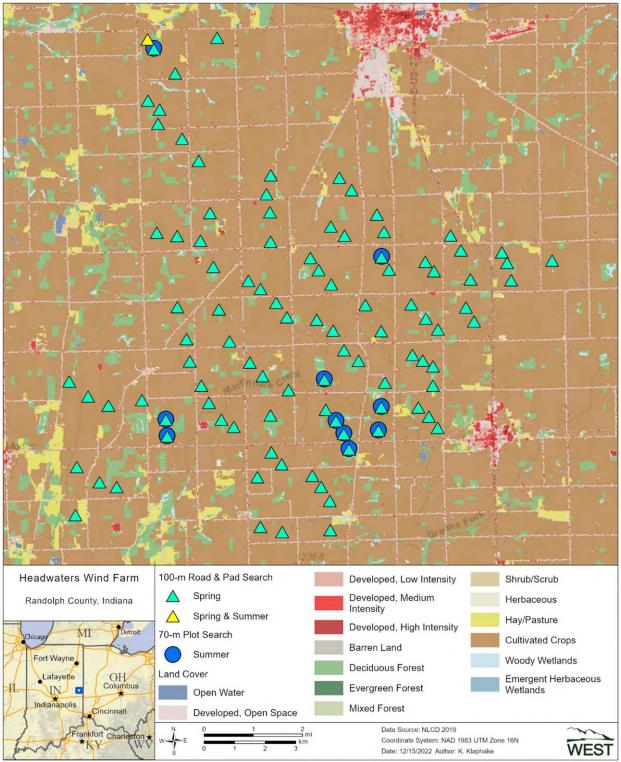


Figure 1. Turbine locations by search type and surrounding land cover at the Headwaters Wind Farm in Randolph County, Indiana.

# METHODS

WEST used Project-specific data from previous post-construction monitoring studies at the Project (2019, 2020, and 2021 studies; see Rodriguez et al. 2020a, 2020b, and McAlexander et al. 2022) and other nearby EDP Renewables operated facilities, to develop a study plan that targeted a *g* of 0.25 to meet the monitoring commitments in the HCP (Headwaters 2019). WEST submitted a study plan to EDPR on January 28, 2022, and received approval from the USFWS on February 23, 2022 (M. Reed, USFWS, pers. comm.).

#### **Standardized Carcass Searches**

#### Number of Turbines Sampled, Search Frequency, and Plot Size

Technicians and dog-handler teams conducted standardized carcass searches (carcass searches) from April 1 – July 31, 2022. Search effort varied by season (Table 3), and was designed to maximize effort when the greatest number of Covered Species were expected to occur.

# Table 3.Search effort by season and plot type at Headwaters Wind Farm in Randolph County,<br/>Indiana, from April 1 – July 31, 2022.

			Number of	
Season	Plot Type	Search Interval	Turbines	Search Team
Spring (April 1–May 15)	100-m road and pad	14 days	100	Technician
Summer (May 16–July 31)	70-m full	3.5 days	10 <sup>1</sup>	Dog-handler Team
	100-m road and pad	3.5 days	1	Technician

m=meter

<sup>1</sup> In 2021, turbine 48 had a summer season Indiana bat fatality. As a result, turbine 48 was reassigned as a summer risk turbine by USFWS and was included as a full plot for the 2022 summer season.

A technician searched the gravel road and pad areas (road and pad plots) under all 100 turbines to a distance of 100 m (328 ft) from the turbine every other week during the spring (Table 3, Figure 2). During the summer, summer risk turbines were searched on a twice-weekly interval. A dog-handler team searched 10 turbines as full plots within a 70-m (230-ft) radius and a technician searched one road and pad plot to a distance of 100 m from the turbine twice a week during the summer (Table 3, Figure 3).

The full plot types for the summer period were originally planned to be a mix of six cleared plots, with the entire search area mowed within 10 to 15 centimeters (four to six inches) in height, and four uncleared plots, with transects mowed in a cross pattern, approximately 1.5 m (4.9 ft) wide, to assist dog teams with plot access but reduce overall crop damages. However, due to dry conditions resulting in poor crop growth, uncleared plots did not require mowing until the final week of surveys, in which three transect lanes were mowed in the soybean (*Glycine max*) crop to assist with plot access. Since plot conditions were consistent throughout a majority of the survey season, analysis did not distinguish between cleared and uncleared plot types. Therefore, this report refers only to full plots.



Figure 2. Representative photo of conditions of a 100-meter road and pad plot at the Headwaters Wind Farm, Randolph County, Indiana, from April 1 – July 31, 2022.

#### Search Methods

WEST used two types of search methods: a technician, or human only visual search, and a doghandler team or olfactory search, where the team consisted of one technician/handler and one trained detection dog. All personnel followed the Project's study plan, including proper handling and reporting of carcasses. Carcass searches were conducted during the day, beginning as early as first light.

#### Road and Pad Searches — Technician

Technicians walked transects spaced five m (16 ft) apart at a rate of approximately 45–60 meters per minute (m/min; 148–197 ft/min) on all gravel road and pad areas within 100 m of the turbine. The technicians scanned the area for carcasses on both sides of the transects out to approximately 2.5 m (8.2 ft) to ensure full visual coverage of each search area. Technician searches were only conducted on road and pad plots.



Figure 3. Representative photo of vegetation conditions in a 70-meter full plot at the Headwaters Wind Farm, Randolph County, Indiana, from April 1 – July 31, 2022.

#### Plot Searches — Dog-handler Team

Dog-handler teams searched 70-m full plots for bat carcasses (Figure 4). Prior to each search, handlers determined the survey start point and the number of transects needed to cover the plot after taking into account wind speed and direction, as well as crop row direction and density (when applicable). Handlers oriented the detection dog to start searches perpendicular to the wind to maximize scent detection. Both wind speed and crop density can affect dispersal of the target odor (i.e., bat carcasses) across the search area. To maximize detection rates during an olfactory search, transect width varied with vegetation density, ranging from five to 10 m (16 to 33 ft) apart in densely vegetated areas, to 10–15 m (33–49-ft) in shorter vegetation. Detection dogs were rewarded with either a food reward or a short play session when they correctly alerted to a bat or bird carcass.



Figure 4. Dog-handler team Noelle Freeman (handler) and Raven (detection dog) performing a full plot search at the Headwaters Wind Farm, Randolph County, Indiana, from April 1 – July 31, 2022.

#### Dog-Handler Team Evaluation

Detection dogs were considered candidates for carcass searches if they met basic temperament and obedience criteria, and demonstrated the trainability to detect bat and/or bird carcasses. Temperament characteristics sought after were high-energy, and a high food or toy drive. Prior to conducting searches at the Project, handlers trained their detection dogs on the scent of bat carcasses following methods derived from search and rescue programs and drug detection (Kay 2012, Helfers 2017). Dogs were initially trained with either cotton scent swabs that had been rubbed on bat carcasses, progressing to dehydrated bats, or directly with dehydrated bat carcasses, at increasing distances over a period of three to four weeks. Once the dog achieved a passing grade of 80% or higher in a scent recognition test, consisting of ten blind trial lineups using dehydrated bats, the dog and handler were evaluated in the field to measure their performance. The detection dog coordinator conducted a two day field evaluation of each doghandler team; after teams achieved a searcher efficiency of 75% or greater for 15–30 dehydrated bats placed during blind evaluation trials, the teams were approved to conduct standardized carcass searches. Because the objective of the study focused on detecting bat carcasses, dogs were not explicitly trained on native bird carcasses; however, all detection dogs alerted on bird carcasses in the field, and handlers rewarded bird finds in the field to encourage future alerts to bird carcasses. A border collie breed was used at the Project as the primary detection dog.

#### Data Collection

Technicians recorded the date, start and end times, technician name, turbine number, type of search and if any carcasses were found for each scheduled search. When a carcass was found, technicians placed a flag near it and continued the search. After searching the entire plot, the technician returned to record information for each carcass on a carcass information data sheet,

including the date and time, species, sex and age (when possible), technician name, turbine number, measured distance from turbine, azimuth from turbine, location of carcass using Geographic Coordinate System (latitude and longitude), habitat surrounding carcass, carcass condition (e.g., intact, scavenged, dismembered), and estimated time of death (e.g., less than one day, two days).

The condition of each carcass found was recorded using the following categories:

- Intact—a carcass that is complete, not badly decomposed, and shows no sign of being fed upon by a predator or scavenger.
- Scavenged—an entire carcass that shows signs of being fed upon by a predator or scavenger, or a portion(s) of a carcass in one location (e.g., wings, skeletal remains, portion of a carcass, etc.), or a carcass that has been heavily infested by insects.
- Dismembered—a carcass found in multiple pieces distributed more than 1.0 m (3.3 ft) apart from one another due to scavenging or other reasons.
- Injured—a bat or bird found alive.

For bird carcasses, the following category was also used:

• Feather spot—Ten or more feathers (excluding down), or two or more primary feathers at one location indicating predation or scavenging of a bird carcass.

Technicians took digital photographs of each carcass, including any visible injuries, and surrounding habitat. No bird carcasses were collected, but a marker was placed next to each bird carcass to avoid duplicate counting. Bat carcasses were collected under the Project's ITP (TE85617C-0), WEST's Federal Native Endangered and Threatened Species Recovery Permit (TE234121-9), and WEST's State Scientific Collection Permit (2229). Technicians placed each bat carcass in a re-sealable plastic bag labeled with a unique carcass identification number, turbine number, and date, for storage in a freezer on site. Leather gloves covered by nitrile or latex gloves were used to handle all bat carcasses to eliminate possible transmission of rabies or other zoonotic diseases, and to reduce possible human scent bias on any carcasses used later in bias trials. Live, injured bats were recorded and considered fatalities for analysis purposes when observed in search areas, and were handled in accordance with permit conditions (left in place).

Carcasses found in non-search areas (e.g., outside of a plot boundary) or outside of the scheduled study period, were recorded as incidental discoveries and documented following the same protocol for those found during standard searches, but were not included in analysis.

#### Carcass Identification and Agency Notification

Identification of bird carcasses was verified by biologists with significant field experience in identification of birds and their feathers. A federally permitted bat biologist (TE33467D-0)

identified all bat carcasses via photographs throughout the survey period, or in hand at the end of the surveys. The USFWS and the Indiana Department of Natural Resources (IDNR) would have been notified within 24 hours of the positive identification of any state- or federally listed species, but none were identified during the searches.

Tissue samples were collected from heavily scavenged or decomposed carcasses that could not be positively identified and had potential to be a Covered Species. Tissue samples were submitted to a USFWS-approved laboratory, the East Stroudsburg University Wildlife Genetics Institute for identification. Bat carcasses that were heavily scavenged but did not have potential to be a Covered Species (i.e., fur was present on the wing, or the forearms measured more than 42 millimeters [1.6 inch] long) were identified to the closest genus or group possible and were not sent off for further identification. Bat carcasses, or representative hair/tissue samples from individual carcasses, are to be delivered to USFWS by January 31, 2023.

#### **Bias Trials**

#### Searcher Efficiency Trials

The objective of the searcher efficiency trials was to estimate the probability that a carcass was found by searchers. Searcher efficiency trials were conducted in the same areas where carcass searches occurred. Technicians conducting carcass surveys did not know when searcher efficiency trials were being conducted or the location of the trial carcasses. Trial carcasses consisted of eastern red bats (*Lasiurus borealis*), silver-haired bats (*Lasionycteris noctivagans*), and big brown bats (*Eptesicus fuscus*) that had previously been found on site or provided by Indiana State University. Sixty-seven carcasses were placed across all seasons and plot types to account for differences in search conditions due to vegetation, topography, or weather.

Multiple trials were conducted in each season to measure potential changes in plot conditions on searcher efficiency over time. Each trial carcass was discreetly marked with a black zip-tie and/or a piece of electrical tape around the upper forelimb for identification as a study carcass. Carcasses were dropped from waist-height or higher and allowed to land in a random posture. The trial administrator walked in a meandering path and dropped trial carcasses for detection dogs the day prior to the next search to allow time for the scent to pool and disperse, and to eliminate a direct scent trail. For technician search trials, the trial administrator placed carcasses prior to the technician searching the plot, either the night before or the morning of searches depending on work schedules.

Searchers had one chance to locate trial carcasses during the first search after carcass placement. The number and location of trial carcasses found during the subsequent search were recorded, and the number of trial carcasses available for detection during each search was determined immediately after each trial by the person responsible for distributing the carcasses. Following searches, any carcasses that were not detected were checked to confirm availability. Forty-five trial carcasses were left in place and used for carcass persistence trials.

#### Carcass Persistence Trials

The objective of carcass persistence trials was to estimate the average probability a carcass would persist, or be available for detection in the field, given the search interval. Carcasses could be removed by scavenging or rendered undetectable by typical farming activities. A minimum of 15 trial carcasses were placed in each season and plot type to incorporate the effects of varying weather and scavenger densities on carcass persistence. No more than two trial carcasses were placed in a plot to avoid potential over-seeding and attracting scavengers.

Technicians monitored the trial carcasses over a 30-day period according to the following schedule, as closely as possible. Carcasses were checked daily for the first four days, then on day 7, 10, 14, 20, and 30. Trial carcasses were monitored until they were completely removed or the trial period ended. Dog-handler teams were used on the 70-m full plots to determine when carcasses were removed, while technicians determined the status of carcasses placed on 100-m roads and pad plots.

#### Search Area Mapping

Plot boundaries were used to verify if carcasses were found inside the search areas, and to inform the distribution of carcasses around turbines to estimate the number of carcasses that fell inside or outside of search areas. Full plots and road and pad plots were delineated in prior survey years using a 72-m radius GIS projection and Eos sub-meter global positioning satellite unit to designate non-searchable areas within the plots. An additional 2.0 m (6.6 ft) were added to the radius of the plots to account for the width of the turbine tower.

#### **Quality Assurance and Quality Control**

Quality assurance and quality control (QA/QC) measures were implemented at all stages of the study, including in the field, during data entry and analysis, and report writing. Following field surveys, technicians were responsible for inspecting data forms for completeness, accuracy, and legibility. Potentially erroneous data were identified using a series of database queries. Irregular codes or data suspected as questionable were discussed with the technician and/or Project manager. Errors, omissions, or problems identified in later stages of analysis were traced back to the raw data forms, and appropriate changes and measures were implemented. A Microsoft® SQL Server database was developed to store, organize, and retrieve survey data. All data forms and electronic data files were retained for reference.

#### **Statistical Analysis**

The Evidence of Absence (EoA; Dalthorp et al. 2017) modeling framework was used to estimate take of Covered Species. EoA was used with data collected in the field to estimate the overall probability of detecting a bat carcass, the take rate of Covered Species, and the number of Covered Species fatalities that occurred. Data used in the EoA model included number of Covered Species fatalities, fatality spatial data from all bats found during surveys, the results of searcher efficiency and carcass persistence trials, the seasonal arrival distribution of bats (described below), and the detection reduction factor (k; described below).

#### Searcher Efficiency Estimation

Searcher efficiency was estimated separately for technicians and dog-handler teams to account for different modes of detection (i.e., technicians use sight, whereas dogs use scent). EoA uses raw searcher efficiency data (e.g. number of found and available trial carcasses) to inform overall probability of detection. However, to determine if searcher efficiency data should be pooled, or separated by strata such as season and/or plot type, searcher efficiency was modeled using logistic regression. For both technicians and dog-handler team models, model selection was completed using an information theoretic approach known as AICc, or corrected Akaike Information Criterion (Burnham and Anderson 2002). The best model was selected as the most parsimonious model within two AICc units of the model with the lowest AICc value. Searcher efficiency data were input into the EoA software according to the model selection results.

The change in searcher efficiency between successive searches was defined by a parameter called the detection reduction factor (k) that can range from zero to one. When k is zero, it implies a carcass missed on the first search would never be found on subsequent searches. A k of one implies searcher efficiency remains constant no matter how many times a carcass was missed. Huso et al. (2017) estimated a value of k = 0.67 for bats, and this value was used to calculate bat fatality estimates using EoA, per the HCP.

#### Carcass Persistence Rate Estimation

Data collected during carcass persistence trials were used to estimate the probability carcasses remained available to be located by the searcher, given the search interval (i.e., the time between scheduled searches). The average probability a carcass persisted was estimated using an interval-censored survival regression with four potential distributions: exponential, log-logistic, lognormal, and Weibull distributions (Kalbfleisch and Prentice 2002, Dalthorp et al. 2018). Season and plot type were used as potential covariates. The best model was selected as the most parsimonious model within two AICc units of the model with the lowest AICc value. The parameter estimates of the selected model ( $\alpha$  [shape] and  $\beta$  [scale], including the 95% Cl of  $\beta$ ) were used as inputs in the EoA Single Class module.

#### Area Adjustment

The search area adjustment accounted for unsearched areas beneath turbines, and was calculated as a probability that ranged from zero to one. The area adjustment was estimated as the product of the proportion of searched area around each turbine and a carcass-density distribution. A truncated weighted maximum likelihood (TWL) modeling approach (Khokan et al. 2013) was used to estimate the carcass-density distribution using site-specific fatality locations. The TWL approach uses weights based on probability of detection and the proportion of area searched in each 1.0-m annulus around the turbine. Distributions considered were normal, gamma, Gompertz, Rayleigh and Weibull (parameterized according to R Development Core Team [2016] and Yee [2015]). The best model was selected using AlCc. The proportion of area searched was calculated in a Geographic Information System as the amount of area searched divided by the total area searched at each 1.0-m annulus around the turbine.

#### Carcasses Excluded from Analysis

Carcasses were excluded from analysis when the carcass was discovered outside of the spatial and temporal scope of the survey design. For example, carcasses found outside a designated plot were not included in the analysis because the TWL fitting procedure accounts for unsearched areas. Carcasses found prior to the start of surveys (e.g., a carcass found on a plot in the spring that was estimated to have died prior to April 1) were also excluded because the carcass occurred outside of the study period. Note that carcasses found in a plot incidentally (e.g. found by maintenance personnel) were included in the analysis if that plot had a scheduled search in the future. If a carcass of a Covered Species had been found outside of the spatial or temporal scope of the survey design it would still be excluded from the area correction estimate, but would be included in the EoA fatality estimate following Dalthorp et al. 2020.

#### Covered Species Take and Detection Probability Estimates

EoA was used to estimate the detection probability (*g*) median cumulative take to-date ( $M^*$ ), mean annual take rate ( $\lambda$ ), and evaluate the probability that the estimated take rate ( $\lambda$ ) exceeded the expected take rate ( $\tau$ ) for Indiana bat and northern long-eared bat (i.e., Covered Species). Estimates were calculated using the EoA method (Dalthorp et al. 2017), using the Single Class, Multiple Class, and Multiple Years modules of EoA.

The EoA Single Class module is used to estimate the detection probability for each independent search stratum (e.g. season by plot type combination). This resulted in alpha ( $\alpha$ ) and beta ( $\beta$ ) parameters that defined the beta distribution of detection probability in each stratum. The EoA Multiple Class module was then used to combine detection probability distributions across strata (i.e, 70-m full plots, and road and pad plots), with weights for each class (density-weighted proportion, or "DWP" in the software) defined by the within-season sampling fraction, relative turbine operations, and seasonal risk. The Multiple Class module of EoA requires weights (DWP) to sum to 1.0 (representing 100% of the risk to bats). When this module is used, unsearched strata are represented with near-zero detection probabilities and beta distribution parameters are set to Ba = 0.01 and Bb = 1,000 (a detection probability of  $10^{-5}$ ) for unsearched areas within each stratum. The results from the Multiple Years module (Ba and Bb parameters for the detection probability for the permit term to date) were used to estimate  $M^*$  (the median cumulative take over the life of the permit),  $\lambda$  (the underlying annual take rate across the single monitoring period) and its 95% CI, and the probability that  $\lambda > \tau$ , where  $\tau$  is the authorized take number divided by the number of years in the permit. Appendix E shows how the compliance metrics were calculated using the EoA Graphical User Interface<sup>2</sup>. For this study, cross-season relative turbine operations were calculated as the number of visits in each season, during which turbines were operating, divided by the total number of visits in each season. Values were re-scaled to sum to one across the three seasons. Weights for spring, summer, and fall were based on the carcass arrival proportions from the Midwest Wind Energy Multi-Species HCP (USFWS 2016): 0.07% in spring, 36% in summer, and 57% in fall. However, the present analysis only covered a portion of the spring season. Consequently, spring and summer arrival proportions were rescaled to 0.16 and

<sup>&</sup>lt;sup>2</sup> There may be minor differences between screen shots and the results in the main text because EoA is a stochastic estimator, leading to slightly different estimates each time the modules are run.

0.84 because the Year 4 study was conducted to complete the initial three-year HCP monitoring requirement (which began with a partial summer and fall Year 1 study), and thus fall was not included in the 2022 estimate. The summer risk estimate was multiplied by 0.11 to reflect relatively fewer risk turbines during summer, and the seasonal weights were again rescaled to sum to 1.0 for spring and summer (Table 4). These values defined the weights for combing the Beta distribution parameters across seasons.

Furthermore, the Multiple Years Module was used to estimate the site-wide, cumulative detection probability for 2022. The EoA Multiple Years Module requires the input  $\rho$ , which weights the years appropriately for combining Beta distribution parameters. In this module, the weights ( $\rho$  in the software) need not sum to 1.0, and a weight of 1.0 is assumed to represent a typical risk year. Weights may be more or less than 1.0 based on turbine operations that differ between years. The 2022 detection probability was combined with the 2019 – 2021 detection probabilities using the multiple years module. The EoA Multiple Years module requires the input  $\rho$ , which weights the years appropriately. Annual weights for 2019 -2021 are described in Rodriguez et al. 2020a, 2020b and McAlexander et al. 2022. The value for  $\rho$  was set to 0.88 for 2019 because the ITP was issued part way through summer, meaning about 88% of total annual risk was realized in 2019. In 2020, the Project was fully operational for all three seasons, so  $\rho$  was set to 1. In 2021, there were comparatively long periods of turbine down-time across the facility, and as described in that year's report,  $\rho$  was set to 0.91. The value for  $\rho$  in 2022 was 0.16 because only spring and summer seasons were studied (Appendix D4).

Table 4.	Seasonal arrival proportions for the fall season at Headwaters Wind Farm in	
	Randolph County, Indiana, from April 1 – July 31, 2022.	

Season	Seasonal Arrival Proportion	Cross-season Relative Operation	Re-Scaled Season Weights
Spring (April 1 – May 15)	0.16	1	0.64
Summer (May 16 – July 31)	0.84	.11	0.36

The results from the Multiple Years module (Ba and Bb parameters for the detection probability for the permit term to date) were used to estimate  $M^*$  (the median cumulative take over the life of the permit),  $\lambda$  (the underlying annual take rate over the past four monitoring periods) and its 90% CI, and the probability that  $\lambda > \tau$ , where  $\tau$  is the authorized take number divided by the number of years in the permit. Appendix D shows how the compliance metrics were calculated using the EoA Graphical User Interface.

#### Adaptive Management Triggers

The estimates from the EoA analysis were used to test two adaptive management triggers: a short-term test of whether the estimated take rate exceeded the expected take rate and a long-term test of whether permitted take had been met (Dalthorp and Huso 2015). Both the short- and long-term triggers were tested individually for each of the Covered Species.

#### Evidence of Absence Short-term Trigger

The EoA short-term trigger was designed as an early warning signal that the project may be on the path to exceeding permitted take (T) by the end of the permit term. The short-term trigger was

designed to determine if an adaptive management response was needed to prevent the cumulative take estimate from actuating a response to the long-term trigger test. The short-term trigger tests if the estimated annual take rate ( $\lambda$ ) exceeded the expected take rate ( $\tau = T \div$  years in permit) at a confidence level of  $\alpha = 0.05$ , per the HCP. The Project short-term trigger was designed to evaluate a rolling window of six years of post-construction monitoring data. If, within any 6-year rolling window, the estimated take rate exceeds the expected take rate with 95% confidence, the short-term trigger would be met, indicating that the minimization plan in the HCP may need to be adjusted to ensure that the median cumulative take estimate (M\*) remains within the permitted limit over the ITP term. Four survey periods of data were used in this analysis, 2019 – 2022.

#### Evidence of Absence Long-term Trigger

The EoA long-term trigger was designed to test if the cumulative take to date was equal to or greater than the permitted take (T). Per the HCP, cumulative take to date ( $M^*$ ) was estimated at a confidence level of  $\alpha = 0.5$  (using the median, or 50<sup>th</sup> credible bound, of the posterior distribution of estimated mortality). If the cumulative take to date at  $\alpha = 0.5$  was less than or equal to the total permitted take ( $M^* \le T$ ), then no changes are necessary. If the cumulative take to date at  $\alpha = 0.5$  was greater than the total permitted take ( $M^* > T$ ) then the take limit was met and the Project must enact avoidance measures.

# RESULTS

# Standardized Carcass Searches

Five hundred twenty-six searches were conducted during the spring and summer monitoring seasons; 13 searches (less than 3%) were missed due to turbine maintenance, weather constraints, and/or safety hazards.

No federally or state-listed bat or bird species were found. Sixty bat carcasses and 15 bird carcasses were found during standardized searches and incidentally (Appendix A). The most commonly found bat species were eastern red bat (28 carcasses; 46.67%) and silver-haired bat (19 carcasses; 31.7%), followed by hoary bat (*Aestus cinereus*; seven carcasses; 11.7%) and big brown bat (six carcasses; 10.0%).

#### **Statistical Analysis**

#### Bias Trials

# Searcher Efficiency Trials

Sixty-seven bats were placed for searcher efficiency trials on eleven separate dates, of which 53 remained available for search teams to find across all plot types. Because dog-handler teams only searched full plots in the summer season, and all full plots were considered a single stratum, there were no covariates to model. The best-fit model for searcher efficiency on road and pad plots did not support the inclusion of season as a covariate, meaning there was not a statistically

meaningful difference between searcher efficiency rates across seasons (Appendix B1, B2). Searcher efficiency rates ranged from 81% on road and pad plots to 84% on full plots.

#### Carcass Persistence Trials

Forty-five carcasses were placed to estimate carcass persistence. The best-fit model for carcass persistence rates was a lognormal distribution with no covariates. Carcass persistence did not vary by season or plot type (Appendix B3, B4). The average probability that a carcass persisted through a 14-day search interval was 0.49 (90% CI: 0.39-0.59) for road and pad plots in spring (Figure 5). The probability of a carcass persisting through a 3.5-day search interval was 0.68 (90% CI: 0.57-0.78) on road and pad plots and full plots in summer (Figure 5).

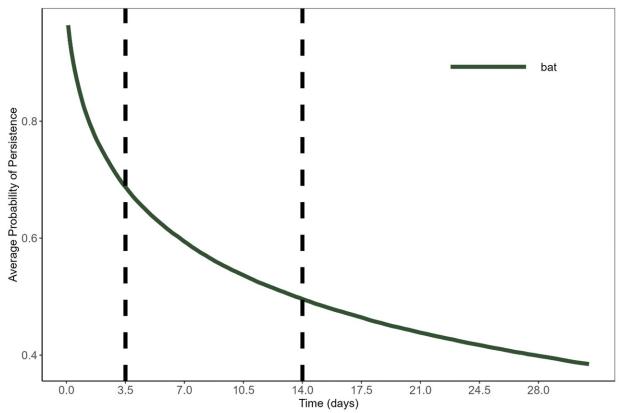


Figure 5. The average probability of persistence, in days, at different search intervals at Headwaters Wind Farm, Randolph County, Indiana, from April 1 – July 31, 2022.

#### Area Adjustment

Nine of the 60 bats found during the monitoring season were excluded from modeling the area adjustment for EoA. Four bat carcasses were excluded from analysis because they were found off plot. Another five bats were excluded because their estimated time of death was prior to the start of surveys (Appendix C1).

The best-fit model for the carcass-density distribution of bats (Figure 6) with respect to distance from turbine base was a Weibull model (Appendix C2). Area adjustments were calculated

separately for bats on the 100-m road and pad plots and 70-m full plots. Carcass counts in the spring were not large enough to model carcass distributions separately by season and were pooled across seasons to inform one overall carcass-density distribution. The area adjustment for bats at 100-m road and pad plots was 0.09 in spring and 0.07 for summer; the area adjustment for bats at 70-m full plots was 0.96 (Appendix C3).

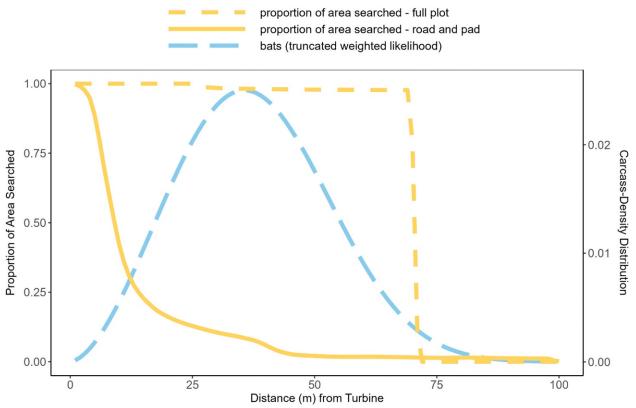


Figure 6. Density of bat carcasses per area searched at road and pad plots and full plots at Headwaters Wind Farm, Randolph County, Indiana, from April 1 – July 31, 2022.

#### Covered Species Take Estimates

No Covered Species carcasses were found during the 2022 study. To date, seven Indiana bats and zero northern long-eared bats have been found under the ITP. The probability of detection distribution (g) achieved for the 2022 monitoring period had a mean of 0.22 (90% CI: 0.19–0.25; Table 5). Inputs required to run the EoA Single Class module and stratum-specific g distribution values and inputs required for the Multiple Class module are described in Appendix D with representative inputs into EoA Software v2.0.7 provided in Appendix E.

	•				
Year	Ba <sup>1</sup>	Bb <sup>1</sup>	$P^2$	g	90% Cl <sup>3</sup>
2019 (fall only)	26.57	97.63	0.88	0.21	0.16-0.28
2020	522.96	1728.37	1	0.23	0.22-0.25
2021	304.41	576.54	0.91	0.34	0.32-0.37
2022 (spring + summer)	98.54	348.59	0.16	0.22	0.19-0.25
Short-Term Trigger (Based on total 3 Year monitoring requirements))	326.09	923.17	NA	0.26	0.24-0.28
Long-Term Trigger (Cumulative)	326.09	923.17	NA	0.26	0.24-0.28

Table 5.Study period and overall probabilities of detection (g), Ba, and Bb and  $\rho$  for the<br/>Headwaters Wind Farm in Randolph County, Indiana, from 2019 – 2022.

<sup>1</sup> Ba and Bb are the parameters for the Beta distribution used to characterize the probability of detection. The *g* value is the mean of that distribution.

 $^2\,\rho$  is the weight in the weighted average that was used to combine the probability of detection distributions across years.

<sup>3</sup>CI = confidence interval

Mean annual take rates based on the complete initial three years of HCP-required monitoring (2019–2022) were estimated to be 9.61 (90% CI: 4.63-16.09) Indiana bats per year and 0.64 (90% CI: 0.00 - 2.46) northern long-eared bats per year (Table 6). The expected average annual take rates reported in the HCP were 9.55 Indiana bats per year and 2.53 northern long-eared bats per year (Headwaters 2019).

Cumulative take under the ITP to-date (2019–2022),  $M^*$ , at  $\alpha = 0.5$  (50<sup>th</sup> credible bound), was estimated to be 27 Indiana bats and zero northern long-eared bats (Table 7). The total take permitted by the ITP is 258 Indiana bats and 68 northern long-eared bats over the 27-year permit term (Table 7; Headwaters 2019).

#### Adaptive Management Triggers

#### Evidence of Absence Short-term Trigger

The short-term trigger assesses the probability that the estimated take rate exceeded the expected take rate,  $Pr(\lambda > \tau)$ . At a 95% confidence level ( $\alpha = 0.05$ ),  $Pr(\lambda > \tau)$  must be greater than or equal to 0.95 for the short-term trigger to fire. For Indiana bat,  $Pr(\lambda > \tau) = 0.46$  and northern long-eared bat,  $Pr(\lambda > \tau) = 0.05$  (Table 6). Neither probability meets or exceeds 0.95, indicating the short-term trigger was not met and no adaptive management actions are necessary (Table 6; Figure 7; Headwaters 2019).

Table 6.	Probability the estimated take rates exceeded the expected take rates for studies
	conducted within the rolling average interval at the Headwaters Wind Farm, Randolph
	County, Indiana, ITP Years 1-4 (2019 – 2022).

		Short-Term Trigger		
Species	Mean λ (90% Cl)	Rate ( $\tau$ )	<b>Pr(λ &gt;</b> τ) <sup>1</sup>	Fires at $\alpha = 0.05$ ?
Indiana bat	9.61 (4.63-16.09)	9.55	0.46	No
Northern long-eared bat	0.64 (0.00–2.46)	2.53	0.05	No

<sup>1</sup> Pr( $\lambda > \tau$ ) reads, "the probability that  $\lambda$  (the annual take rate) is greater than  $\tau$  (the expected annual take rate based on the total permitted take, used as a threshold for adaptive management)." If this probability was less than 0.95 (e.g.,  $\alpha = 0.05$  for a 1-sided test), then no adaptive management was triggered because there was not sufficient evidence that the estimated annual take rate was greater than the expected annual take rate.

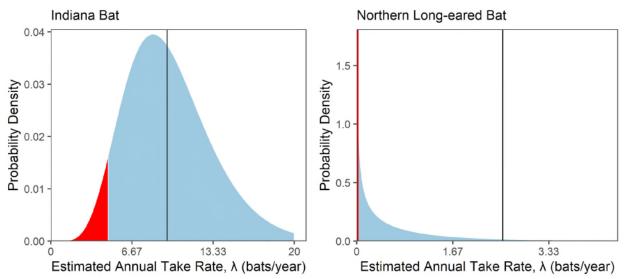


Figure 7. Estimated annual take rate ( $\lambda$ ) bats per year at Headwaters Wind Farm, Randolph County, Indiana, ITP Years 1-4 (2019 – 2022).

Note: The red region of the posterior distributions shows the region of the lower 5% quantile of the distributions (red region may not be visible when the posterior distribution is skewed heavily toward zero). The vertical line marks the expected take rate. The short-term trigger evaluates whether the vertical line falls within or to the left of the red region of the posterior distributions. For both species, the short-term trigger was not met because the vertical line (expected take rate) was not within or to the left of the red regions. In other words, the probability that estimated take rate was greater than the expected take rate does not exceed 95%.

#### Evidence of Absence Long-term Trigger

The estimated cumulative take to date,  $M^*$  at  $\alpha = 0.5$  (50<sup>th</sup> credible bound), was below the total permitted take for both Covered Species (Table 7). The long-term trigger was not met and the Project is in compliance for both species because  $M^* < T$  for both species. Therefore, an avoidance response is not necessary.

	within the ITP term to date at Headwaters Wind Farm, Randolph County, Indiana, fro ITP Years 1-4 (2019 – 2022).							
Species		Cumulative take (M*)	Permitted take	Long-term trigger fires at $\alpha = 0.5$ ?				
Species		(171)	(1)	mes al 0 – 0.5?				

27

0

258

68

No

No

Table 7.	Cumulative take estimate to date using Evidence of Absence for studies conducted
	within the ITP term to date at Headwaters Wind Farm, Randolph County, Indiana, from
	ITP Years 1-4 (2019 – 2022).

# CONCLUSIONS

Indiana bat (50<sup>th</sup> credible bound)

northern long-eared bat (50<sup>th</sup> credible bound)

The post-construction monitoring effort completed in 2022 was consistent with the HCP's monitoring requirements and the Project's 2022 study plan. No Covered Species carcasses were found despite a high probability of detection in 2022. Estimates of potential take for the Covered Species were below the levels authorized by the ITP and no adaptive management was necessary.

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Appendix A. Carcasses found during the 2022 Post-Construction Monitoring Surveys

		Distance from Turbine				Physical	Aided
Found Date	Common Name	(m)	Turbine	Search Type	Plot Type	Condition	Search
Bats							
04/04/2022	silver-haired bat	3	22	carcass search	twice per month road and pad	scavenged	no
04/06/2022	eastern red bat	10	92	incidental	twice per month road and pad	scavenged	no
05/02/2022	silver-haired bat	21	20	carcass search	twice per month road and pad	scavenged	no
05/05/2022	silver-haired bat	40	72	carcass search	twice per month road and pad	intact	no
05/12/2022	eastern red bat	53	77	incidental	twice per week cleared plot	scavenged	no
05/16/2022	eastern red bat	90	48	carcass search <sup>2</sup>	twice per week cleared plot	scavenged	yes <sup>1</sup>
05/16/2022	silver-haired bat	70	48	carcass search	twice per week cleared plot	dismembered	yes <sup>1</sup>
05/16/2022	silver-haired bat	38	48	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
05/19/2022	silver-haired bat	69	48	carcass search	twice per week cleared plot	dismembered	yes <sup>1</sup>
05/20/2022	eastern red bat	39	39	carcass search	twice per week cleared plot	intact	yes1
05/24/2022	hoary bat	64	77	carcass search	twice per week cleared plot	scavenged	yes1
06/02/2022	big brown bat	44	24	carcass search	twice per week cleared plot	scavenged	yes1
06/02/2022	big brown bat	55	48	carcass search	twice per week cleared plot	scavenged	yes1
06/02/2022	eastern red bat	50	71	carcass search	twice per week cleared plot	scavenged	yes1
06/02/2022	eastern red bat	53	71	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/02/2022	eastern red bat	69	71	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/02/2022	silver-haired bat	90	48	carcass search <sup>2</sup>	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/06/2022	eastern red bat	12	24	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/06/2022	eastern red bat	56	48	carcass search	twice per week cleared plot	intact	yes <sup>1</sup>
06/07/2022	eastern red bat	5	95	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/09/2022	silver-haired bat	23	41	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/10/2022	silver-haired bat	37	40	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/10/2022	silver-haired bat	38	96	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/13/2022	big brown bat	22	79	incidental <sup>2</sup>	n/a	intact	no

Appendix A. Carcasses found at the Headwaters Wind Farm, Randolph County, Indiana, from April 1 – July 31, 2022.

		Distance from Turbine				Physical	Aided
Found Date 06/13/2022	Common Name hoary bat	<u>(m)</u> 43	Turbine 41	Search Type carcass search	Plot Type twice per week cleared plot	Condition scavenged	Search yes <sup>1</sup>
06/14/2022	eastern red bat	43 34	77			-	•
				carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/14/2022	silver-haired bat	24	77	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/14/2022	silver-haired bat	25	95	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/16/2022	eastern red bat	18	70	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/16/2022	silver-haired bat	21	24	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/16/2022	silver-haired bat	18	70	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/17/2022	eastern red bat	57	39	carcass search	twice per week cleared plot	scavenged	yes1
06/17/2022	eastern red bat	63	39	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/17/2022	hoary bat	68	95	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/17/2022	hoary bat	9	96	carcass search	twice per week cleared plot	dismembered	yes <sup>1</sup>
06/17/2022	silver-haired bat	67	96	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/21/2022	eastern red bat	59	24	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/21/2022	eastern red bat	31	24	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/22/2022	eastern red bat	59	96	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/22/2022	silver-haired bat	36	40	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/22/2022	silver-haired bat	14	40	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/22/2022	silver-haired bat	65	40	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/23/2022	silver-haired bat	59	48	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/24/2022	hoary bat	32	39	carcass search	twice per week cleared plot	intact	yes <sup>1</sup>
06/28/2022	eastern red bat	58	39	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/28/2022	eastern red bat	43	39	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/28/2022	eastern red bat	17	95	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
07/04/2022	big brown bat	29	70	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
07/08/2022	eastern red bat	42	39	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
						0	

Appendix A. Carcasses found at the Headwaters Wind Farm, Randolph County, Indiana, from April 1 – July 31, 2022.

	-	Distance from Turbine	-	-		Physical	Aided
Found Date	Common Name	(m)	Turbine	Search Type	Plot Type	Condition	Search
07/08/2022	eastern red bat	54	95	carcass search	twice per week cleared plot	scavenged	yes1
07/10/2022	eastern red bat	40	24	carcass search	twice per week cleared plot	dismembered	yes <sup>1</sup>
07/10/2022	eastern red bat	33	48	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
07/11/2022	eastern red bat	38	39	incidental	twice per week cleared plot	scavenged	yes <sup>1</sup>
07/11/2022	hoary bat	44	96	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
07/16/2022	eastern red bat	23	40	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
07/21/2022	eastern red bat	27	24	carcass search	twice per week cleared plot	scavenged	yes1
07/28/2022	hoary bat	51	24	carcass search	twice per week cleared plot	intact	yes1
07/29/2022	big brown bat	21	96	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
07/29/2022	eastern red bat	32	77	carcass search	twice per week cleared plot	scavenged	yes1
08/16/2022	big brown bat	9	74	incidental <sup>2</sup>	n/a	intact	no
Birds							
04/04/2022	horned lark	5	13	carcass search	twice per month road and pad	scavenged	no
04/04/2022	horned lark	17	23	carcass search	twice per month road and pad	intact	no
04/07/2022	American kestrel	8	100	incidental	twice per month road and pad	intact	no
05/02/2022	pine warbler	95	18	carcass search	twice per month road and pad	intact	no
05/12/2022	horned lark	48	77	incidental	twice per week cleared plot	scavenged	no
05/13/2022	red-tailed hawk	12	25	incidental	twice per week road and pad	scavenged	no
05/16/2022	chestnut-sided warbler	82	48	carcass search <sup>2</sup>	twice per week cleared plot	scavenged	yes <sup>1</sup>
05/16/2022	unidentified passerine	102	48	carcass search <sup>2</sup>	twice per week cleared plot	dismembered	yes1
05/19/2022	red-winged blackbird	2	70	carcass search	twice per week cleared plot	dismembered	yes <sup>1</sup>
05/20/2022	red-tailed hawk	68	77	carcass search	twice per week cleared plot	dismembered	yes <sup>1</sup>
05/23/2022	red-eyed vireo	79	71	carcass search <sup>2</sup>	twice per week cleared plot	scavenged	yes <sup>1</sup>

Appendix A. Carcasses found at the Headwaters Wind Farm, Randolph County, Indiana, from April 1 – July 31, 2022.

		Distance from Turbine				Physical	Aided
Found Date	Common Name	(m)	Turbine	Search Type	Plot Type	Condition	Search
05/24/2022	red-eyed vireo	40	40	carcass search	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/02/2022	yellow-billed cuckoo	90	48	carcass search <sup>2</sup>	twice per week cleared plot	scavenged	yes <sup>1</sup>
06/13/2022	horned lark	22	22	incidental <sup>2</sup>	n/a	scavenged	no
07/04/2022	purple martin	73	24	carcass search <sup>2</sup>	twice per week cleared plot	intact	yes <sup>1</sup>

Appendix A. Carcasses found at the Headwaters Wind Farm, Randolph County, Indiana, from April 1 – July 31, 2022.

<sup>1</sup> dog aided search

<sup>2</sup> Carcass was found outside the search area and/or outside study period

Appendix B. Searcher Efficiency and Carcass Persistence Model Fitting Results

Appendix B1. Searcher	efficiency m	odels for	70-meter	plots a	it the	Headwaters	Wind	Farm,
Randolph County	, Indiana from	n April 1, 2	2022 to July	y 31, 202	22 (n=	21).		

Covariates	k Value	AICc	Delta AICc
No covariates	0.67	22.66	0 <sup>1</sup>

<sup>1</sup> Selected model.

AICc = Corrected Akaike Information Criterion.

Delta AICc = Change in AICc

Appendix B2. Searcher efficiency models for road and pad plots at the Headwaters Wind Farm, Randolph County, Indiana from April 1, 2022 to July 31, 2022 (n = 32).

Covariates	k Value	AICc	Delta AICc
No Covariates	0.67	29.87	0 <sup>1</sup>
Season	0.67	32.15	2.28

<sup>1</sup> Selected model.

AICc = Corrected Akaike Information Criterion.

Delta AICc = Change in AICc

Appendix B3. Carcass persistence models with covariates and distributions for bats at the Headwaters Wind Farm, Randolph County, Indiana from April 1, 2022 to July 31, 2022 (n = 45).

= 49):				
Location Covariates	Scale Covariates	Distribution	AICc	Delta AICc
Season	No Covariates	lognormal	185.72	0
Season	No Covariates	loglogistic	185.78	0.06
Season	Season	loglogistic	185.87	0.15
Season	Season	lognormal	185.96	0.24
PlotSearchType	No Covariates	lognormal	186.12	0.40
PlotSearchType	No Covariates	loglogistic	186.26	0.54
Season	No Covariates	Weibull	186.33	0.61
Season + PlotSearchType	No Covariates	lognormal	186.96	1.24
Season + PlotSearchType	No Covariates	loglogistic	187.12	1.40
Season + PlotSearchType	Season	lognormal	187.43	1.71
Season + PlotSearchType	Season	loglogistic	187.44	1.72
Season + PlotSearchType	No Covariates	Weibull	187.48	1.76
No Covariates	No Covariates	lognormal	187.59	1.87 <sup>1</sup>
No Covariates	No Covariates	loglogistic	187.61	1.89
PlotSearchType	No Covariates	Weibull	187.66	1.94
PlotSearchType	PlotSearchType	lognormal	187.89	2.17
PlotSearchType	PlotSearchType	loglogistic	188.04	2.32
No Covariates	Season	lognormal	188.05	2.33
No Covariates	Season	loglogistic	188.10	2.38
Season	Season	Weibull	188.15	2.43
Season	Season + PlotSearchType	loglogistic	188.40	2.68

Season	Season + PlotSearchType	lognormal	188.50	2.78
PlotSearchType	Season + PlotSearchType	lognormal	188.58	2.86
Season + PlotSearchType	PlotSearchType	lognormal	188.69	2.97
PlotSearchType	Season + PlotSearchType	loglogistic	188.74	3.02
Season + PlotSearchType	PlotSearchType	loglogistic	188.82	3.10
No Covariates No Covariates	PlotSearchType PlotSearchType	lognormal loglogistic	188.90 188.92	3.18 3.20
Season + PlotSearchType	Season	Weibull	189.48	3.76
Season + PlotSearchType	PlotSearchType	Weibull	189.82	4.10
No Covariates	No Covariates	Weibull	189.92	4.20
PlotSearchType	PlotSearchType	Weibull	189.97	4.25
Season + PlotSearchType	Season + PlotSearchType	lognormal	190.04	4.32
Season + PlotSearchType	Season + PlotSearchType	loglogistic	190.05	4.33
No Covariates	Season + PlotSearchType	lognormal	190.18	4.46
No Covariates	Season + PlotSearchType	loglogistic	190.23	4.51
Season	Season + PlotSearchType	Weibull	190.67	4.95
PlotSearchType	Season + PlotSearchType	Weibull	191.72	6.00
No Covariates	Season	Weibull	191.99	6.27
No Covariates	PlotSearchType	Weibull	192.01	6.29
Season + PlotSearchType	Season + PlotSearchType	Weibull	192.14	6.42
No Covariates	Season + PlotSearchType	Weibull	194.31	8.59
Season	-	exponential	210.73	25.01
Season + PlotSearchType	-	exponential	211.25	25.53
PlotSearchType	-	exponential	214.50	28.78
No Covariates	-	exponential	219.79	34.07
10111				

<sup>1</sup> Selected model

AICc = Corrected Akaike Information Criterion

Delta AICc = Change in AICc

# Appendix B4. Carcass persistence top model with covariates, distributions, and model parameters for the Headwaters Wind Farm, Randolph County, Indiana, from April 1 – July 31, 2022.

	-	Estimated Median Removal	-	-
Size Class	Distribution	Times (days)	Parameter 1	Parameter 2
Bat	Lognormal <sup>1</sup>	5.10	meanlog = 1.63	sdlog = 2.592

<sup>1</sup> Parameterization follows the base R parameterization for this distribution.

Appendix C. Truncated Weighted Likelihood (TWL) Area Adjustment Model Fitting Results

Appendix C1. Number and percent (%) of bat carcasses found and total included in the area
adjustment calculation for the Headwaters Wind Farm, Randolph County, Indiana, from
April 1 – July 31, 2022.

		d in Area stment		e Search ea¹		e Study iod¹	То	otal
Species	Total	%	Total	%	Total	%	Total	%
eastern red bat	25	49.02	1	25.00	2	40.00	28	46.67
silver-haired bat	15	29.41	1	25.00	3	60.00	19	31.67
hoary bat	7	13.73	0	0	0	0	7	11.67
big brown bat	4	7.84	2	50.00	0	0	6	10.00
Total	51	100	4	100	5	100	60	100

<sup>1</sup> Carcasses outside of search area or study period not included in analysis

Appendix C2. Search area adjustment models for bats from the Headwaters Wind Farm, Randolph	
County, Indiana, from April 1 – July 31, 2022.	

Distribution	AICc	Delta AICc	
Weibull	1,076.85	0 <sup>1</sup>	
normal	1,078.83	1.98	
gamma	1,084.39	7.54	
Gompertz	1,085.73	8.88	
Rayleigh	1,089.02	12.17	

<sup>1</sup> Selected model

## Appendix C3. Truncated weighted maximum likelihood search area adjustment estimates for the Headwaters Wind Farm, Randolph County, Indiana, from April 1 – July 31, 2022 (Bat n=51).

Size	Search Area	-	-	-	-	Area
Class	Туре	Distribution	Season	Parameter 1	Parameter 2	Adjustment
	full plot	Weibull	Summer	2.6655	42.4906	0.96
Bat	road and pad <sup>1</sup>	Weibull	Spring	2.6655	42.4906	0.09
	road and pad <sup>2</sup>	Weibull	Summer	2.6655	42.4906	0.07

<sup>1</sup> All plots were searched as road and pad plots in the spring.

<sup>2</sup> Only one road and pad plot was searched during the summer season.

Appendix D. Inputs for Single Class and Multiple Class Modules in Evidence of Absence

		Search		Spatial	Searcher I	Efficiency	Carcass Pe	rsistence <sup>2</sup>
		Interval	Number of	Coverage	Carcasses	Carcasses	Shape	Scale
Season	Plot Type	(I)	Searches	(a)	Available	Found	(α)	(β)
Spring	100-m road and pad	14	4	0.09	29	25	6.72	1.63
Summer	100-m road and pad	4	21	0.07	29	25	6.72	1.63
Summer	70-m full plot	3.5	22	0.96	24	19	6.72	1.63

Appendix D1. Inputs needed to run Evidence of Absence: Single Class Module for the Headwaters Wind Farm, Randolph County, Indiana, from April 1 – July 31, 2022<sup>1</sup>

<sup>1</sup> k was assumed to equal 0.67 for all strata, per Huso et al. (2017).

<sup>2</sup> A lognormal distribution was used for carcass persistence on both 100-m road and pad plots and 70-m full plots. The 95% upper and lower confidence intervals on β were set to 0.82, 2.44.

m = meter

Appendix D2. Inputs needed to run Evidence of Absence model to combine across plot types within each season: Multiple Class Module for the Headwaters Wind Farm, Randolph County, Indiana, from April 1 – July 31, 2022

Season	Plot Type	Ва	Bb	Within-Season Sampling Fraction
Spring	100-m road and pad	50.02	1186.15	1.0
Summer	100-m road and pad	109.9	2303.49	0.09
Summer	70-m full plot	40.65	28.29	0.91

m = meter

Appendix D3. Inputs needed to run Evidence of Absence model to combine across seasons: Multiple Class Module for the Headwaters Wind Farm, Randolph County, Indiana, from April 1 – July 31, 2022

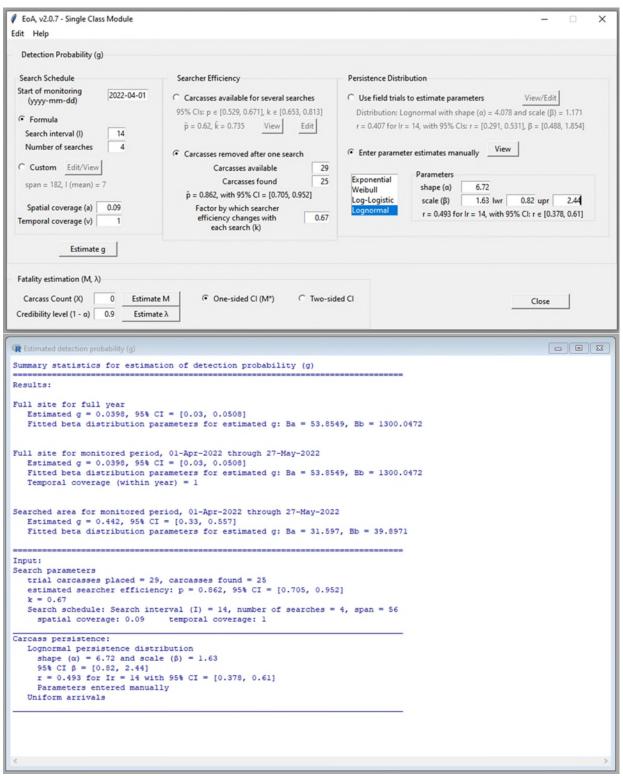
Season	Ba	Bb	Weights (DWP <sup>1</sup> )
Spring (April 1–May 15)	50.02	1186.15	0.64
Summer (May 16–July 31)	46.39	39.48	0.36

<sup>1</sup> DWP = Density-weighted proportion

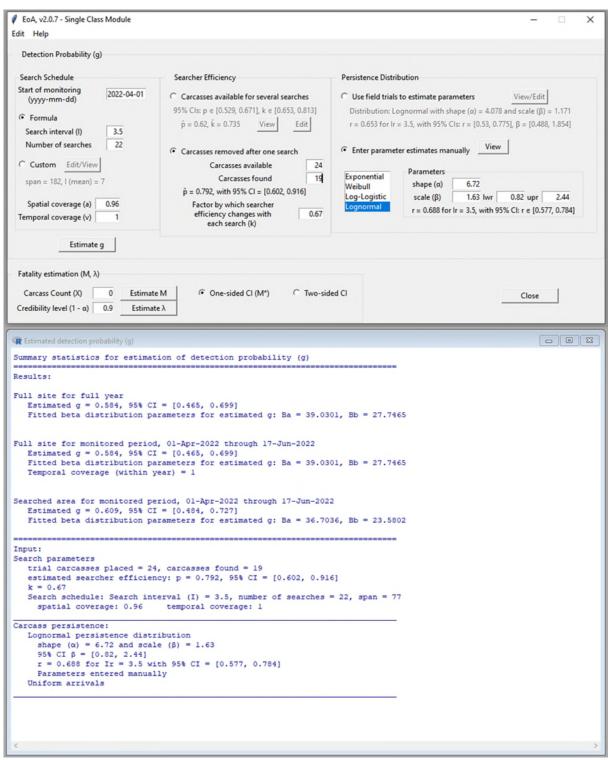
Appendix D4.	Inputs needed to run Evidence of Absence model to combine across years:
Multiple Y	Years Module for the Headwaters Wind Farm in Randolph County, Indiana, from
2019 – 202	22.

Year	Ba	Bb	Weights (ρ)
2019	26.57	97.63	0.88
2020	522.96	1728.37	1.0
2021	304.41	576.54	0.91
2022	98.54	348.59	0.16

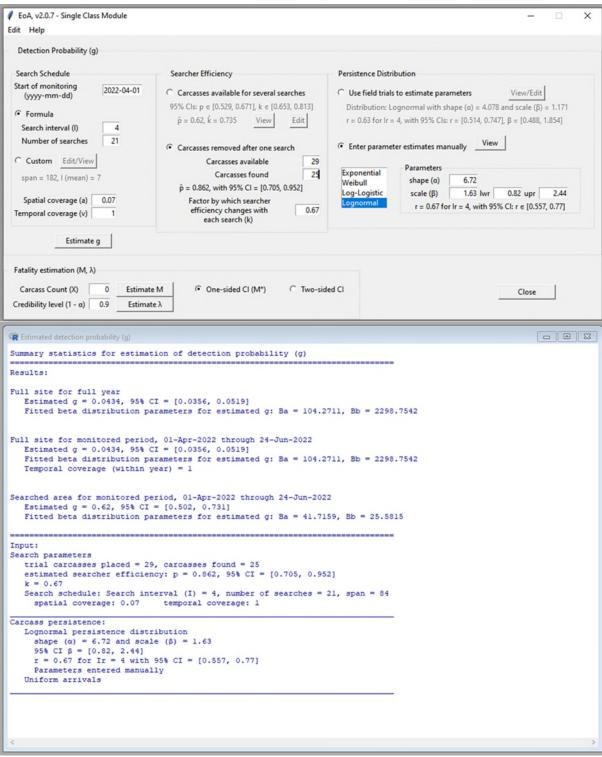
Appendix E. Screenshots of Inputs for Single Class and Multiple Class Modules in Evidence of Absence



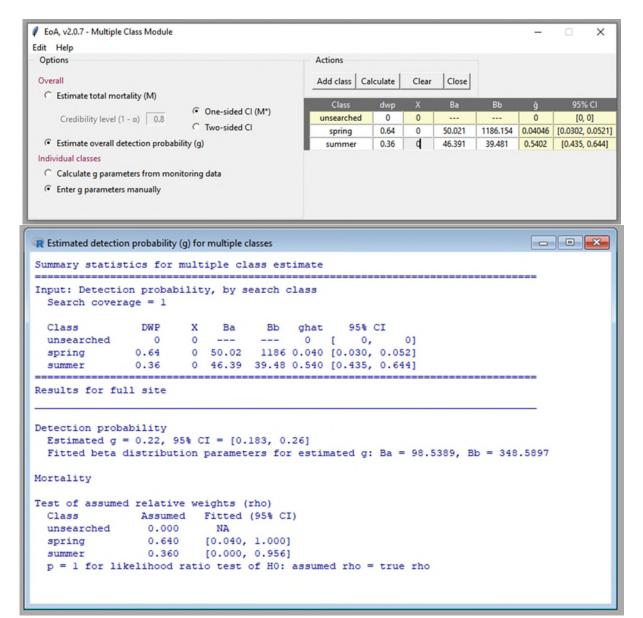
Appendix E1. Screen shot of Evidence of Absence (v2.0.7) graphical user interface, Single Class Module inputs for Spring 2022, 100-meter road and pad searches at 100 turbines with a blade length of 55 meters, searched at a 14-day interval.



Appendix E2. Screen shot of Evidence of Absence (v2.0.7) graphical user interface, Single Class Module inputs for Summer 2022, 70-meter full plot searches at 11 turbines with a blade length of 55 meters, searched at a 3.5-day interval.



Appendix E3. Screen shot of Evidence of Absence (v2.0.7) graphical user interface, Single Class Module inputs for Summer 2022, 100-meter road and pad searches at 1 turbine with a blade length of 55 meters, searched at a 3.5-day interval.



Appendix E4. Screen shot of Evidence of Absence (v2.0.7) graphical user interface, Multiple Class Module inputs and output for Spring and Summer 2022, (n=100 in spring, 11 in summer), searched at a 14-day interval in the spring and a 3.5-day interval in the summer.

	– 🗆 X
EoA, v2.0.7 - Multiple Years Module	
Edit Help	Options
Past monitoring and operations data	
Year p X Ba Bb ĝ 95% Cl	Fatalities
See legend 1.04 1 37.14 135.67 0.2149 [0.157, 0.279]	Estimate M Credibility level (1 - a) 0.5
2020 1 1 522.96 1728.37 0.2323 [0.215, 0.25]	Total mortality     One-sided CI (M*)
2021 0.91 5 304.41 576.54 0.3455 [0.315, 0.377]	C Two-sided Cl
	Project parameters Total years in project 3
	Total years in project  Mortality threshold (T)  154
	C Track past mortality
	C Projection of future mortality and estimates
	Future monitoring and operations
	G g and p unchanged from most recent year
	G g and p constant, different from most recent year
	g 0.08 95% CI: 0.07 0.09 p 1
	C g and p vary among future years
	Average Rate
	C Estimate average annual fatality rate (λ)
	Annual rate theshold (t) 9.55
	C Credibility level for Cl (1-a) 0.9
	( Short-term rate ( $\lambda > \tau$ ) Term: 3 $\alpha$ 0.05
	C Reversion test $(\lambda < \rho \tau)$ $\rho$ 0.6 $\alpha$ 0.1
	Actions
	Calculate Close
R Mortality over 3 years	
Summary statistics for total mortality through 3 years	^
Results $M^* = 27$ for $1 - \alpha = 0.5$ , i.e., $P(M \le 27) \ge 50^{\frac{1}{2}}$	
$M^* = 27$ for $1 - \alpha = 0.5$ , i.e., $P(M \le 27) >= 50$ %	
M* = 27 for 1 - $\alpha$ = 0.5, i.e., P(M <= 27) >= 50% Estimated overall detection probability: g = 0.261, 95% CI = [0.237,	0.286]
M* = 27 for 1 - α = 0.5, i.e., P(M <= 27) >= 50% Estimated overall detection probability: g = 0.261, 95% CI = [0.237, Ba = 326.52, Bb = 924.04	
M* = 27 for 1 - $\alpha$ = 0.5, i.e., P(M <= 27) >= 50% Estimated overall detection probability: g = 0.261, 95% CI = [0.237,	
$ \begin{split} M^* &= 27 \mbox{ for } 1 - \alpha = 0.5, \mbox{ i.e., } P(M <= 27) >= 50  \\ \mbox{Estimated overall detection probability: } g = 0.261, \mbox{ 95  CI = [0.237,  Ba = 326.52, \mbox{ Bb = 924.04} \\ \mbox{Estimated baseline fatality rate: lambda = 9.769, \mbox{ 95  CI = [4.05, \mbox{ 18}] \\  \end{array} $	1
$M^* = 27$ for $1 - \alpha = 0.5$ , i.e., $P(M <= 27) >= 50$ % Estimated overall detection probability: $g = 0.261$ , 95% CI = [0.237, Ba = 326.52, Bb = 924.04 Estimated baseline fatality rate: lambda = 9.769, 95% CI = [4.05, 18 Test of assumed relative weights (rho) and potential bias	
$M^*=27$ for $1-\alpha=0.5,$ i.e., $P(M<=27)>=504$ Estimated overall detection probability: $g=0.261,\;954$ CI = [0.237, Ba = 326.52, Bb = 924.04 Estimated baseline fatality rate: lambda = 9.769, 954 CI = [4.05, 18 Test of assumed relative weights (rho) and potential bias Assumed rho = 956 CI = 1.04 [0.062, 1.636]	1
$ \begin{split} M^* &= 27 \mbox{ for } 1 - \alpha = 0.5, \mbox{ i.e., } P(M <= 27) >= 50  \\ \mbox{Estimated overall detection probability: } g = 0.261, 95 \mbox{ CI} = [0.237, \mbox{ Ba} = 326.52, \mbox{ Bb} = 924.04 \\ \mbox{Estimated baseline fatality rate: lambda} = 9.769, 95 \mbox{ CI} = [4.05, 18] \\ \mbox{Test of assumed relative weights (rho) and potential bias} \\ \mbox{Assumed rho} = 95 \mbox{ CI} \\ \mbox{ lo.066, l.661]} \\ \mbox{ 1 0.056, l.601]} \end{split} $	1
$ \begin{split} M^* &= 27 \mbox{ for } 1 - \alpha = 0.5, \mbox{ i.e., } P(M <= 27) >= 50  8 \\  \mbox{Estimated overall detection probability: } g = 0.261, 95 \mbox{ CI} = [0.237, Ba = 326.52, Bb = 924.04 \\  \mbox{Estimated baseline fatality rate: lambda = 9.769, 95 \mbox{ CI} = [4.05, 18] \\  \mbox{Test of assumed relative weights (rho) and potential bias} \\ \mbox{Assumed rho} \qquad 95 \mbox{ CI} \\ \mbox{ l.0.04} \mbox{ [0.062, l.636] } \\ \mbox{ 1 } \mbox{ [0.066, l.601] } \\ \mbox{ 0.91} \mbox{ [0.684, 2.588]} \end{split} $	1
$ \begin{split} M^* &= 27 \mbox{ for } 1 - \alpha = 0.5, \mbox{ i.e., } P(M <= 27) >= 50  \\ \mbox{Estimated overall detection probability: } g = 0.261, 95 \mbox{ CI} = [0.237, \mbox{ Ba} = 326.52, \mbox{ Bb} = 924.04 \\ \mbox{Estimated baseline fatality rate: lambda} = 9.769, 95 \mbox{ CI} = [4.05, 18] \\ \mbox{Test of assumed relative weights (rho) and potential bias} \\ \mbox{Assumed rho} = 95 \mbox{ CI} \\ \mbox{ lo.066, l.661]} \\ \mbox{ 1 0.056, l.601]} \end{split} $	1
$ \begin{split} \mathbf{M}^{*} &= 27 \text{ for } 1 - \alpha = 0.5, \text{ i.e., } P(M <= 27) >= 50\$ \\ \text{Estimated overall detection probability: } g = 0.261, 95\$ CI = [0.237, Ba = 326.52, Bb = 524.04 \\ \text{Estimated baseline fatality rate: lambda = 9.769, 95\$ CI = [4.05, 18] \\ \text{Test of assumed relative weights (rho) and potential bias} \\ \text{Assumed tho} \qquad 55\$ CI \\ 1.04 \qquad [0.062, 1.636] \\ 1 \qquad [0.062, 1.636] \\ 0.91 \qquad [0.664, 2.588] \\ p = 0.26589 \text{ for likelihood ratio test of H0: assumed rho = true rho} \\ \text{Quick test of relative bias: 1.129} \end{split} $	1
$ \begin{split} M^* &= 27 \mbox{ for } 1 - \alpha = 0.5, \mbox{ i.e., } P(M <= 27) >= 50  \\ \mbox{Estimated overall detection probability: } g = 0.261, 95 \mbox{ CI} = [0.237, Ba = 326.52, Bb = 924.04 \\ \mbox{Estimated baseline fatality rate: lambda = 9.769, 95 \mbox{ CI} = [4.05, 18] \\ \mbox{Test of assumed relative weights (rho) and potential bias} \\ \mbox{Assumed rho} \qquad 95 \mbox{ CI} \\ \mbox{ loc.} 0.062, \mbox{ l.636]} \\ \mbox{ 1 } [0.062, \mbox{ l.636]} \\ \mbox{ 0.91 } [0.664, \mbox{ 2.588]} \\ \mbox{ p = 0.26589 for likelihood ratio test of H0: assumed rho = true rho} \end{split} $	1
$ \begin{split} M^{*} &= 27 \mbox{ for } 1 - \alpha = 0.5, \mbox{ i.e., } P(M <= 27) >= 50  8 \\  \mbox{Estimated overall detection probability: } g = 0.261, 95 \mbox{ CI} = [0.237, Ba = 326.52, Bb = 924.04 \\  \mbox{Estimated baseline fatality rate: lambda = 9.769, 95 \mbox{ CI} = [4.05, 18 \\  \mbox{Test of assumed relative weights (rho) and potential bias} \\  \mbox{Assumed rho} \qquad 95 \mbox{ CI} \\  \mbox{ locs, l.636]} \\ \mbox{ 1 } [0.062, l.636] \\ \mbox{ 0.91 } [0.664, 2.588] \\ \mbox{p = 0.26589 for likelihood ratio test of H0: assumed rho = true rho} \\  \mbox{Quick test of relative bias: l.129} \\ \mbox{Posterior distribution of M} \\ \mbox{ m } p(M = m) p(M > m) \\ \mbox{ 0 } 0.0000 \\ \mbox{ 1 } 0.000 \\ \end{tabular}$	1
M = 27 for 1 - α = 0.5, i.e., P(M <= 27) >= 50% Estimated overall detection probability: g = 0.261, 95% CI = [0.237, Ba = 326.52, Bb = 924.04 Estimated baseline fatality rate: lambda = 9.769, 95% CI = [4.05, 18 Test of assumed relative weights (rho) and potential bias Assumed rho 95% CI = [0.236] 1.04 [0.056, 1.601] 0.91 [0.684, 2.588] p = 0.26589 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 1.129 Posterior distribution of M m p(M = m) p(M > m) 0 0.0000 1.0000	1
M = 27 for 1 - α = 0.5, i.e., P(M <= 27) >= 50% Estimated overall detection probability: g = 0.261, 95% CI = [0.237, Ba = 326.52, Bb = 924.04 Estimated baseline fatality rate: lambda = 9.769, 95% CI = [4.05, 18 Test of assumed relative weights (rho) and potential bias Assumed rho 95% CI 1.04 [0.062, 1.636] 1 [0.056, 1.601] 0.91 [0.664, 2.588] p = 0.26589 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 1.129 Posterior distribution of M m p(M = m) p(M > m) 0 0.0000 1.0000 1.0000 2 0.0000 1.0000	1
M = 27 for 1 - α = 0.5, i.e., P(M <= 27) >= 50% Estimated overall detection probability: g = 0.261, 95% CI = [0.237, Ba = 326.52, Bb = 924.04 Estimated baseline fatality rate: lambda = 9.769, 95% CI = [4.05, 18 Test of assumed relative weights (rho) and potential bias Assumed rho 95% CI 1.04 [0.062, 1.636] 1 [0.056, 1.601] 0.91 [0.684, 2.588] p = 0.26589 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 1.129 Posterior distribution of M m p(M = m) p(M > m) 0 0.0000 1.0000 2 0.0000 1.0000 3 0.0000 1.0000	1
M = 27 for 1 - α = 0.5, i.e., P(M <= 27) >= 50% Estimated overall detection probability: g = 0.261, 95% CI = [0.237, Ba = 326.52, Bb = 924.04 Estimated baseline fatality rate: lambda = 9.769, 95% CI = [4.05, 18 Test of assumed relative weights (rho) and potential bias Assumed rho 95% CI 1.04 [0.062, 1.636] 1 [0.056, 1.601] 0.91 [0.684, 2.588] p = 0.26589 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 1.129 Posterior distribution of M m p(M = m) p(M > m) 0 0.0000 1.0000 2 0.0000 1.0000 3 0.0000 1.0000 5 0.0000 1.0000 6 0.0000 1.0000	1
$ \begin{split} M^* = 27 \mbox{ for } 1 - \alpha = 0.5, \mbox{ i.e., } P(M <= 27) >= 50 \mbox{ Stimated overall detection probability: } g = 0.261, 95 \mbox{ CI} = [0.237, Ba = 326.52, Bb = 924.04 \\  \mbox{ Estimated baseline fatality rate: lambda = 9.769, 95 \mbox{ CI} = [4.05, 18 \\  \mbox{ Test of assumed relative weights (rho) and potential bias } \\  \mbox{ Assumed rho} \qquad 95 \mbox{ CI} \\ 1.04 & [0.062, 1.636] \\ 1 & [0.065, 1.601] \\ 0.91 & [0.664, 2.588] \\ \mbox{p} = 0.26589 \mbox{ for likelihood ratio test of H0: assumed rho = true rho} \\ \mbox{ Quick test of relative bias: 1.129 } \\ \mbox{ Posterior distribution of M } \\ \mbox{m} & p(M = m) p(M > m) \\ 0 & 0.0000 & 1.0000 \\ 1 & 0.0000 & 1.0000 \\ 3 & 0.0000 & 1.0000 \\ 3 & 0.0000 & 1.0000 \\ 5 & 0.0000 & 1.0000 \\ 6 & 0.0000 & 1.0000 \\ \end{array} $	1
$\begin{array}{llllllllllllllllllllllllllllllllllll$	1
$\begin{array}{l} M^{*}=27 \mbox{ for }1-\alpha=0.5, \mbox{ i.e., }P(M<=27)>=508\\ \\ \mbox{Estimated overall detection probability: }g=0.261, 958 \mbox{ CI}=[0.237, Ba=326.52, Bb=924.04\\ \\ \mbox{Estimated baseline fatality rate: lambda}=9.769, 958 \mbox{ CI}=[4.05, 18]\\ \\ \mbox{Test of assumed relative weights (rho) and potential bias}\\ \\ \mbox{Assumed rho} & 958 \mbox{ CI}\\ 1.04 & [0.062, 1.636]\\ 1 & [0.056, 1.601]\\ 0.91 & [0.664, 2.588]\\ \\ \mbox{p}=0.26589 \mbox{ for likelihood ratio test of H0: assumed rho}= true rho\\ \\ \\ \mbox{Quick test of relative bias: 1.129}\\ \\ \\ \mbox{Posterior distribution of M}\\ \\ \mbox{m} & p(M=m) \mbox{ p}(M>m)\\ 0 & 0.0000 & 1.0000\\ \\ \mbox{l} & 0.0000 & 1.0000\\ \\ \mbox{l} & 0.0000 & 1.0000\\ \\ \mbox{d} & 0.0000 & 1.000\\ \\ \mbox{d} & $	1
$\begin{array}{llllllllllllllllllllllllllllllllllll$	1
$\begin{array}{l} M^{*}=27 \mbox{ for }1-\alpha=0.5, \mbox{ i.e., }P(M<=27)>=508\\ \\ \mbox{Estimated overall detection probability: }g=0.261, 958 \mbox{ CI}=[0.237, \mbox{ Ba}=326.52, \mbox{ Bb}=924.04\\ \\ \mbox{Estimated baseline fatality rate: lambda}=9.769, 958 \mbox{ CI}=[4.05, 18]\\ \\ \mbox{Test of assumed relative weights (rho) and potential bias}\\ \\ \mbox{Assumed rho}=958 \mbox{ CI}\\ 1.04 \ [0.062, 1.636]\\ 1 \ [0.056, 1.601]\\ 0.91 \ [0.684, 2.588]\\ \\ \mbox{p}=0.26589 \mbox{ for likelihood ratio test of H0: assumed rho}= true rho\\ \\ \\ \mbox{Quick test of relative bias: l.129}\\ \\ \\ \mbox{Posterior distribution of M}\\ \\ \mbox{m} \ p(M=m) \ p(M>m)\\ 0 \ 0.0000 \ 1.0000\\ \\ 2 \ 0.0000 \ 1.0000\\ \\ 3 \ 0.0000 \ 1.0000\\ \\ 5 \ 0.0000 \ 1.0000\\ \\ 6 \ 0.0000 \ 1.0000\\ \\ 7 \ 0.0000 \ 1.0000\\ \\ 8 \ 0.0020 \ 0.9997\\ \\ 9 \ 0.0018 \ 0.9972\\ \\ 11 \ 0.0034 \ 0.9938\\ \\ 12 \ 0.0088 \ 0.9938\\ \\ 13 \ 0.0088 \ 0.9792 \\ \\ \end{array}$	1
$\begin{array}{llllllllllllllllllllllllllllllllllll$	1
$\begin{array}{llllllllllllllllllllllllllllllllllll$	1
$\begin{array}{l} M^*=27 \mbox{ for } 1-\alpha=0.5, \mbox{ i.e., } P(M<=27)>=508\\ \\ \mbox{Estimated overall detection probability: } g=0.261, 958 \mbox{ CI}=[0.237, Ba=326.52, Bb=924.04\\ \\ \mbox{Estimated baseline fatality rate: lambda}=9.769, 958 \mbox{ CI}=[4.05, 18]\\ \\ \mbox{Test of assumed relative weights (rho) and potential bias}\\ \\ \mbox{Assumed rho} & 958 \mbox{ CI}\\ 1.04 & [0.062, 1.636]\\ 1 & [0.056, 1.601]\\ 0.91 & [0.664, 2.588]\\ \\ \mbox{p}=0.26589 \mbox{ for likelihood ratio test of H0: assumed rho}= true rho\\ \\ \\ \mbox{Quick test of relative bias: 1.129}\\ \\ \\ \mbox{Posterior distribution of M}\\ \\ \mbox{m} & p(M=m) \mbox{ p(M > m)}\\ 0 & 0.0000 & 1.0000\\ \\ \mbox{l} & 0.0000 & 1.0000\\ \\ \mbox{l} & 0.0000 & 1.0000\\ \\ \mbox{d} & 0.0000\\ \\ \mbox{d} & 0.000$	1

## Appendix E5. Multiple years M\* Indiana bat.

\*Due to the limitations of the EoA GUI with 2 study periods that don't encompass a full monitoring year, we combined the 2019 and 2022 study periods into one full monitoring year. Rho values of 0.88 in 2019 and 0.16 in 2022 were summed for a combined rho value of 1.04 for this stratum. We used these 3 full study periods to estimate lambda and M\*, as well as to evaluate the short term and long term triggers. The alpha and beta parameters for combined 2019 and 2022 monitoring periods were obtained by using the multiple class module to combine g values. The multiple class module requires the weights to sum to 1.0, so the 0.88 and 1.06 values were rescaled by dividing by 1.04.

🖉 EoA, v2.0.7 - Multiple Years Module	X
Edit Help	^
con ricp	Options
Past monitoring and operations data	Fatalities
Year p X Ba Bb ĝ 95% Cl	C Estimate M Credibility level (1 - a) 0.5
See legend 1.04 1 37.14 135.67 0.2149 [0.157, 0.279]	One-sided CI (M*)
2020         1         1         522.96         1728.37         0.2323         [0.215, 0.25]           2021         0.91         5         304.41         576.54         0.3455         [0.315, 0.377]	Total mortality     Two-sided Cl
	Project parameters
	Total years in project 3
	Mortality threshold (T) 154
	C Track past mortality
	C Projection of future mortality and estimates Future monitoring and operations
	g and p unchanged from most recent year
	C g and p constant, different from most recent year
	g 0.08 95% Cl: 0.07 0.09 p 1
	C g and p vary among future years
	Average Rate <sup>(*)</sup> Estimate average annual fatality rate (λ)
	Annual rate theshold (t) 9.55
	C Credibility level for Cl (1-a) 0.9
	(• Short-term rate $(\lambda > \tau)$ Term: 3 $\alpha$ 0.05
	C Reversion test ( $\lambda < \rho \tau$ ) $\rho = 0.6 \alpha = 0.1$
	Actions
	Calculate Close
R Short-term Trigger	
Short-term trigger: Test of average fatality rate (lambda) over 3 ye Years: See legend - 2021	613
Results	
<pre>Estimated overall detection probability: g = 0.261, 95% CI = [0.237, Ba = 326.52, Bb = 924.04</pre>	0.286]
Estimated annual fatality rate over the past 3 years: lambda = 9.607 P(lambda > 9.55) = 0.4566	7, 95% CI = [3.99, 17.7]
Compliance: Cannot infer lambda > 9.55 with 95% credibility	
Input	
Threshold for short-term rate (tau) = 9.55 per year	
Period rel_wt X Ba Bb ghat 95% CI	
See legend         1.040         1         37.14         135.7         0.215         [0.157, 0.279]           2020         1.000         1         523         1728         0.232         [0.215, 0.250]	
2021 0.910 5 304.4 576.5 0.346 [0.315, 0.377]	
¢.	>

Appendix E6. Lambda Indiana bat

EoA, v2.0.7 - Multiple Years Module	
	– 🗆 X
Edit Help	Options
Bud wer in the second	
Past monitoring and operations data Year p X Ba Bb ĝ 95% Cl	Fatalities
Year         ρ         X         Ba         Bb         ĝ         95% Cl           See legend         1.04         0         37.14         135.67         0.2149         [0.157, 0.279]	Festimate M Credibility level (1 - α) 0.5
2020 1 0 522.96 1728.37 0.2323 [0.215, 0.25]	Total mortality     One-sided CI (M*)
2021 0.91 d 304.41 576.54 0.3455 [0.315, 0.377]	C Two-sided Cl
	Project parameters
	Total years in project
	Mortality threshold (T) 154
	C Track past mortality
	C Projection of future mortality and estimates
	Future monitoring and operations
	G and p unchanged from most recent year
	G and p constant, different from most recent year
	g 0.08 95% Cl: 0.07 0.09 p 1
	C g and p vary among future years
	Average Rate
	C Estimate average annual fatality rate (λ)
	Annual rate theshold (t) 2.53
	C Credibility level for CI (1-a) 0.9
	(* Short-term rate ( $\lambda > \tau$ ) Term: 3 $\alpha$ 0.05
	C Reversion test ( $\lambda < \rho \tau$ ) $\rho$ 0.6 $\alpha$ 0.1
	Actions
	Calculate Close
R Mortality over 3 years	- • •
	^
Summary statistics for total mortality through 3 years	
Results	
$M^* = 0$ for $1 - \alpha = 0.5$ , i.e., $P(M <= 0) >= 50$ %	
Estimated overall detection probability: g = 0.261, 95% CI = [0.237, 0.286]	
Ba = 326.52, Bb = 924.04	
Estimated baseline fatality rate: lambda = 0.6512, 95% CI = [0.000645, 3.28]	
	5, 3.28]
Test of assumed relative weights (rho) and potential bias	5, 3.28] Fitted zho
Assumed rho 95% CI	
Assumed rho 95% CI 1.04 [0.009, 2.786]	
Assumed rho 95% CI 1.04 [0.009, 2.786]	
Assumed tho 95% CI 1.04 [0.009, 2.786] 1 [0.011, 2.829] 0.91 [0.006, 2.670]	
Assumed tho 95% CI 1.04 [0.009, 2.786] 1 [0.011, 2.829]	
Assumed rho 95% CI 1.04 [0.009, 2.786] 1 [0.011, 2.829] 0.91 [0.006, 2.670] p = 1 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 0.968	
Assumed rho 95% CT 1.04 [0.009, 2.786] 1 [0.011, 2.829] 0.91 [0.006, 2.670] p = 1 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 0.968 Posterior distribution of M m p(M = m) p(M > m)	
Assumed rho 95% CI 1.04 [0.009, 2.786] 1 [0.011, 2.829] 0.91 [0.006, 2.670] p = 1 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 0.968 Posterior distribution of M m p(M = m) p(M > m) 0 0.5518 0.4482	
Assumed rho 95% CT 1.04 [0.009, 2.786] 1 [0.011, 2.829] 0.91 [0.006, 2.670] p = 1 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 0.968 Posterior distribution of M m p(M = m) p(M > m)	
<pre>Assumed rho 95% CT 1.04 [0.009, 2.786] 1 [0.011, 2.829] 0.91 [0.006, 2.670] p = 1 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 0.968 Posterior distribution of M m p(M = m) p(M &gt; m) 0 0.5518 0.4482 1 0.1689 0.2793 2 0.0958 0.1835 3 0.0557 0.1238</pre>	
Assumed rho 95% CI 1.04 [0.009, 2.786] 1 [0.011, 2.829] 0.91 [0.006, 2.670] p = 1 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 0.968 Posterior distribution of M m p(M = m) p(M > m) 0 0.5518 0.4482 1 0.1689 0.2793 2 0.0585 0.1835 3 0.0587 0.1238 4 0.0389 0.0849 5 0.0260 0.0589	
Assumed rho 95% CT 1.04 [0.009, 2.786] 1 [0.011, 2.829] 0.91 [0.006, 2.670] p = 1 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 0.968 Posterior distribution of M m p(M = m) p(M > m) 0 0.5518 0.4482 1 0.1689 0.2793 2 0.0958 0.1835 3 0.0557 0.1238 4 0.0389 0.0849 5 0.0260 0.0589 6 0.0177 0.0412	
Assumed rho 95% CT 1.04 [0.009, 2.786] 1 [0.011, 2.829] 0.91 [0.006, 2.670] p = 1 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 0.968 Posterior distribution of M m p(M = m) p(M > m) 0 0.5518 0.4482 1 0.1689 0.2793 2 0.0587 0.1238 4 0.0389 0.0849 5 0.0260 0.0589 6 0.0177 0.0412 7 0.0122 0.0290	
Assumed rho 95% CT 1.04 [0.009, 2.786] 1 [0.011, 2.829] 0.91 [0.006, 2.670] p = 1 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 0.968 Posterior distribution of M m p(M = m) p(M > m) 0 0.5518 0.4482 1 0.1659 0.2793 2 0.0958 0.1835 3 0.0557 0.1238 4 0.0359 0.0849 5 0.0260 0.0589 6 0.0177 0.0412 7 0.0122 0.0290 8 0.0085 0.0246	
Assumed rho 95% CT 1.04 [0.009, 2.786] 1 [0.011, 2.829] 0.91 [0.006, 2.670] p = 1 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 0.968 Posterior distribution of M m p(M = m) p(M > m) 0 0.5518 0.4482 1 0.1689 0.2793 2 0.0558 0.1835 3 0.0557 0.1238 4 0.0389 0.0849 5 0.0260 0.0589 6 0.0177 0.0412 7 0.0122 0.0250 8 0.0085 0.0205 9 0.0059 0.0146 10 0.0042 0.0104	
Assumed rho 95% CT 1.04 [0.009, 2.786] 1 [0.011, 2.829] 0.91 [0.006, 2.670] p = 1 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 0.968 Posterior distribution of M m p(M = m) p(M > m) 0 0.5518 0.4482 1 0.1689 0.2793 2 0.0587 0.1238 4 0.0389 0.0849 5 0.0260 0.0589 6 0.0177 0.0412 7 0.0122 0.0290 8 0.0055 0.0265 9 0.0059 0.0146 10 0.0054 0.0074 11 0.0030 0.0074 12 0.0021 0.0053	
Assumed rho 95% CT 1.04 [0.009, 2.786] 1 [0.011, 2.829] 0.91 [0.006, 2.670] p = 1 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 0.968 Posterior distribution of M m p(M = m) p(M > m) 0 0.5518 0.4482 1 0.1689 0.2793 2 0.0595 0.1835 3 0.0597 0.1238 4 0.0389 0.0849 5 0.0260 0.0589 6 0.0177 0.0412 7 0.0122 0.0290 8 0.0085 0.0246 10 0.0055 0.0146 10 0.0030 0.0074 12 0.0021 0.0038	
Assumed rho 95% CT 1.04 [0.009, 2.786] 1 [0.011, 2.829] 0.91 [0.006, 2.670] p = 1 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 0.968 Posterior distribution of M m p(M = m) p(M > m) 0 0.5518 0.4482 1 0.1689 0.2793 2 0.0587 0.1238 4 0.0389 0.0849 5 0.0260 0.0589 6 0.0177 0.0412 7 0.0122 0.0290 8 0.0055 0.0205 9 0.0059 0.0146 10 0.0059 0.0146 10 0.0059 0.0074 12 0.0021 0.0053	
Assumed rho 95% CT 1.04 [0.006, 2.786] 1 [0.011, 2.829] 0.91 [0.006, 2.670] p = 1 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 0.968 Posterior distribution of M m p(M = m) p(M > m) 0 0.5518 0.4482 1 0.1689 0.2793 2 0.0958 0.1835 3 0.0597 0.1238 4 0.0389 0.0849 5 0.0260 0.0589 6 0.0177 0.0412 7 0.0122 0.0290 8 0.00055 0.0205 9 0.0059 0.0146 10 0.0042 0.0104 11 0.0030 0.0074 12 0.0021 0.0053 13 0.0015 0.0038 14 0.0011 0.0028 15 0.0006 0.0020 16 0.0006 0.0020	
Assumed rho 95% CT 1.04 [0.009, 2.786] 1 [0.011, 2.829] 0.91 [0.006, 2.670] p = 1 for likelihood ratio test of H0: assumed rho = true rho Quick test of relative bias: 0.968 Posterior distribution of M m p(M = m) p(M > m) 0 0.5518 0.4482 1 0.1669 0.2793 2 0.0958 0.1835 3 0.0557 0.1238 4 0.0359 0.0849 5 0.0260 0.0589 6 0.0177 0.0412 7 0.0122 0.0290 8 0.0055 0.0205 9 0.0059 0.0146 10 0.0042 0.0104 11 0.0031 0.0053 13 0.0015 0.0038 14 0.0011 0.0028 15 0.0008 0.0020	

## Appendix E7. M\* northern long-eared bat

\*Due to the limitations of the EoA GUI with 2 study periods that don't encompass a full monitoring year, we combined the 2019 and 2022 study periods into one full monitoring year. Rho values of 0.88 in 2019 and 0.16 in 2022 were summed for a combined rho value of 1.04 for this stratum. We used these 3 full study periods to estimate lambda and M\*, as well as to evaluate the short term and long term triggers. The alpha and beta parameters for combined 2019 and 2022 monitoring periods were obtained by using the multiple class module to combine g values. The multiple class module requires the weights to sum to 1.0, so the 0.88 and 1.06 values were rescaled by dividing by 1.04.

EoA, v2.0.7 - Multiple Years Module	- 🗆 X
Edit Help	
	Options
Past monitoring and operations data	Fatalities
Year         ρ         X         Ba         Bb         ĝ         95% Cl           See legend         1.04         0         37.14         135.67         0.2149         [0.157, 0.279]           2020         1         0         522.96         1728.37         0.2323         [0.215, 0.25]           2021         0.91 <b>d</b> 304.41         576.54         0.3455         [0.315, 0.377]	<ul> <li>C Estimate M Credibility level (1 - a) 0.5</li> <li>C Total mortality</li> <li>C One-sided Cl (M*)</li> <li>C Two-sided Cl</li> <li>Project parameters</li> <li>Total years in project 3</li> <li>Mortality threshold (1) 154</li> <li>C Track past mortality</li> <li>C Projection of future mortality and estimates</li> <li>Future monitoring and operations</li> <li>G g and ρ unchanged from most recent year</li> <li>g 0.08 95% Cl: 0.07 0.09 ρ 1</li> <li>C g and ρ vary among future years</li> </ul>
	Average Rate
	Actions Calculate Close
R Short-term Trigger	
Short-term trigger: Test of average fatality rate (lambda) Years: See legend - 2021 Results Estimated overall detection probability: g = 0.261, 95% CI Ba = 326.52, Bb = 924.04	over 3 years
Estimated annual fatality rate over the past 3 years: lambd P(lambda > 2.53) = 0.0469000000000000 Compliance: Cannot infer lambda > 2.53 with 95% credibil	
Input Threshold for short-term rate (tau) = 2.53 per year	
Period         rel_wt         X         Ba         Bb         ghat         95% CI           See legend         1.040         0         37.14         135.7         0.215         [0.157, 0.279]           2020         1.000         0         523         1728         0.232         [0.215, 0.250]           2021         0.910         0         304.4         576.5         0.346         [0.315, 0.377]	

Appendix E8. Lambda northern long-eared bat